

# Effect of Statins on Aneurysmal Subarachnoid Hemorrhage: A Meta-Analysis of Randomized Controlled Trials

Anevrizmal Subaraknoid Kanamaya Statinlerin Etkisi: Randomize Kontrollü Çalışmaların Bir Meta-Analizi

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## ABSTRACT

AIM: The role of statins for treating aneurysmal subarachnoid hemorrhage (aSAH) remains uncertain. In this study, the relevance of different end points was evaluated in order to clarify the action and efficacy of statins.

**MATERIAL and METHODS:** A systematic literature retrieval was carried out to obtain randomized controlled trials (RCTs) from before March 2013 on the use of statins for aSAH. Data extraction and quality evaluation of the studies were performed by 2 investigators. A meta-analysis was performed using Review Manager (RevMan) software version 5.2.3.

**RESULTS:** Seven randomized controlled trials comprising 347 patients that met the inclusion criteria were included in this meta-analysis. Results showed that, in aSAH, statins did not reduce vasospasm on transcranial Doppler (RR=0.80; 95% CI, 0.53-1.21; p=0.29) or improve outcomes (RR=0.92; 95% CI, 0.71-1.20; p=0.54). However, statins were able to decrease delayed ischemic neurological deficits (RR=0.56; 95% CI, 0.41-0.75; p=0.0001) and mortality (RR=0.54; 95% CI, 0.32-0.91; p=0.02) compared with placebo.

**CONCLUSION:** Acute statin treatment might not be a good choice for cerebral vasospasm after aSAH. Further large-scale, well-designed RCTs on this topic are still needed.

KEYWORDS: Statins, Aneurysmal subarachnoid hemorrhage, Meta-analysis, Transcranial doppler

## ÖΖ

AMAÇ: Anevrizmal subaraknoid kanamayı (aSAK) tedavi etmekte statinlerin rolü halen kesin değildir. Bu çalışmada statinlerin etkisi ve etkinliğini açıklığa kavuşturmak için farklı son noktaların önemi değerlendirilmiştir.

YÖNTEM ve GEREÇLER: Mart 2013 öncesinde aSAK için statinlerin kullanımı konusunda randomize kontrollü çalışmaları bulmak için sistematik bir literatür taraması yapılmıştır. Çalışmalardan verilerin alınması ve kalitelerinin değerlendirilmesi 2 araştırmacı tarafından gerçekleştirilmiştir. Review Manager (RevMan) yazılımı versiyon 5.2.3 kullanılarak bir meta-analiz yapılmıştır.

**BULGULAR:** Bu meta analize çalışmaya alma kriterlerini karşılayan 347 hastalık yedi randomize kontrollü çalışma dahil edilmiştir. Sonuçlar aSAK durumunda statinlerin transkraniyal Doppler'de vazospazmı azaltmadığını (RR=0,80; %95 GA, 0,53-1,21; p=0,29) veya sonuçları daha iyi hale getirmediğini (RR=0,92; %95Cl, 0,71-1,20; p=0,54) göstermiştir. Ancak statinler plaseboyla karşılaştırıldığında gecikmiş iskemik nörolojik defisitleri (RR=0,56; %95 GA, 0,41-0,75; p=0,0001) ve mortaliteyi (RR=0,54; %95 GA, 0,32-0,91; p=0,02) azaltmıştır.

SONUÇ: Akut statin tedavisi aSAK sonrasında serebral vazospazm için iyi bir tercih olmayabilir. Bu konuda daha büyük ölçekli ve iyi tasarlanmış randomize kontrollü çalışmalar gereklidir.

ANAHTAR SÖZCÜKLER: Statinler, Anevrizmal subaraknoid kanama, Meta-analiz, Transkraniyal doppler

# INTRODUCTION

Subarachnoid hemorrhage (SAH) is one of the most common cerebrovascular events. Every year, approximately 30,000 people in the United States experience an SAH (28). Early brain injury (27), cerebral vasospasm (CVS) (26), and delayed ischemic neurological deficit (DIND) (12) after aneurysmal SAH (aSAH) strongly determine the patient's prognosis. Many treatment options for preventing these complications are available in clinical practice.

Statins have been administered for aSAH. Atorvastatin can reduce the incidence and severity of CVS and suppress a biomarker of cerebral ischemia after aSAH (20). Simvastatin has a beneficial therapeutic effect on reducing cognitive dysfunction after aSAH (16). Clinical trials have demonstrated that simvastatin reduces CVS and mortality and improves functional outcomes, although these findings were not statistically significant (4). Patients who are treated with statins prior to aSAH tend to have a lower incidence of subsequent CVS (17). However, other studies have shown that simvastatin does not exert a beneficial effect on patients with aSAH (8, 9). Kramer et al. (11) pointed out that if there is a benefit to statin use, it may be smaller than suggested by previous studies.

Thus, the role of statins for treating patients with aSAH remains unclear. The designers of various studies considered whether or not statins reduce the likelihood of CVS and DIND. Therefore, in this analysis, we evaluated the efficacy of statins by comparing different reported outcomes in order to discuss the role of statins for treating aSAH.

# MATERIAL and METHODS

## Inclusion Criteria

First, RCTs only were selected, but blinding and allocation concealment were not initially considered. Second, participants were included according to age (>18 years), a diagnosis of aSAH after emergency hospital admission, and initial aSAH grade according to the World Federation of Neurological Surgeons scale; a computed tomography scan of the head was assessed within 24 hours. All patients underwent endovascular coiling or surgical clipping for aneurysm during the first 72 hours of ictus. After undergoing coiling or clipping for aneurysm, patients were randomly allocated to the statin group or the placebo group. The statin group had received simvastatin or pravastatin, and the placebo group had received placebo or nonstatin treatment for 14 or more days. The primary end points were CVS on transcranial Doppler (TCD), DIND, poor outcome, and mortality. Secondary end points were time to CVS and DIND from ictus, duration of CVS and DIND, and adverse events at follow-up.

# Literature Search

Our literature search followed the methodological guidelines in the *Cochrane Handbook for Systematic Reviews of Interventions* version 5.1.0. Three search engines were used: Medline, EMBASE, and the Cochrane Library, and results were retrieved with following words: Subarachnoid Hemorrhage, statin, simvastatin, pravastatin, and Hydroxymethylglutaryl-CoA Reductase Inhibitors. The search process did not set a limit on results. In addition, the references listed at the end of the studies were manually checked to filter in potentially eligible studies. The last date reached by the search was March 2013.

# Data Extraction and Quality Evaluation

Two investigators (Xu-Dong Ma and Shi-Tao Zhang) independently read titles and abstracts, excluding trials that did not match the inclusion criteria. Patients' baseline characteristics and details of the study designs were extracted from the included studies. Inconsistent data were extracted by a third investigator (Jun-Jie Zhao).

We evaluated the methodological quality of the studies based on criteria described in the *Cochrane Handbook* 5.1.0, including random-sequence generation, allocation concealment, blinding, and incomplete outcome data. In order to describe those aspects of every study, investigators chose the terms "low risk of bias," "high risk of bias," or "unclear" (lacking information or of uncertain bias) to define each study. To clarify ambiguous information about study design or data, investigators contacted authors by e-mail.

#### **Statistical Methods**

For calculating the effective size of dichotomous variables, relative risk (RR) was used. Effective size described with 95% confidence intervals (CIs). Before analyzing trial data, the heterogeneity of the trials was tested using  $l^2$  statistics, a quantitative method for measuring inconsistency across studies. Outcomes with an l<sup>2</sup> statistic of 25%-50% were considered to have low heterogeneity, an l<sup>2</sup> statistic of 50%-75% was linked to moderate heterogeneity, and an  $l^2$ statistic of >75% demonstrated high heterogeneity (5). An l<sup>2</sup> value >50% indicated significant heterogeneity. A fixedeffects model was appropriate for analyzing studies with low heterogeneity (l<sup>2</sup><50%), and a random-effects model was used to analyze the outcome of studies with significant heterogeneity  $(l^2 > 50\%)$  (1). When statistical heterogeneity existed without clinical heterogeneity among various studies, the random-effects model was also used.

Investigators further conducted sensitivity analyses to explore possible explanations for heterogeneity in the overall pooled estimate. For conclusions, a probability value of *P*<0.05 was considered statistically significant for differences. All statistical analyses were performed using Review Manager (RevMan) software version 5.2.3 software (The Nordic Cochrane Centre, The Cochrane Collaboration, Copenhagen, Denmark).

### RESULTS

### **Process for Included Trials**

A total of 95 studies were identified from 3 databases. After reading the titles and abstracts, 78 studies were excluded because they were duplications, reviews, case reports, or animal experiments; the remaining 17 trials focused on statins for treating aSAH. Of these 17 studies included in the meta-analysis, 5 were concerned with the biological effects of statins, 5 were controlled trials that enrolled 930 patients to study clinical efficacy, and 7 were RCTs that enrolled 347 patients and that met the inclusion criteria (2, 4, 7, 13, 14, 23, 25).

#### Characteristics of Included Trials and Quality Evaluation

The main characteristics of the trials included in the metaanalysis are shown in Table 1. Seven studies were published between 2005 and 2013, most of which had small sample size. In 7 of the studies included in this analysis, 2 (7, 23) administered pravastatin or placebo to 178 participates, and 5 (2, 4, 13, 14, 25) administered simvastatin or placebo to 169 patients.

The methodological quality of trials is shown graphically in Figures 1 and 2. Figure 1 shows the risk of bias for included studies, and Figure 2 represents the overall bias of every study aspect. Five trials (2, 4, 13, 23, 25) were randomized, double-blind, placebo-controlled trials; 2 trials (7, 14) were randomized only in the choice of statin or nonstatin treatment, without a blinding method. Four trials (2, 13, 23, 25) included an intention-to-treat analysis.

#### Table I: Characteristics of Trials Included in the Meta-Analysis

First Author (Year of Publication)	No. of Patients (Statin/ Placebo Group)	Patient Characteristics	Statin Group	Placebo Group Regimen	Definitions of CVS
Tseng (2005) (23)	80 (40/40)	Patients with aSAH (18-84 years old); clipping/coiling was performed in 27/7 in the statin group and 25/6 in the placebo group	Capsules containing pravastatin 40 mg/d were commenced within 72 h of ictus and continued for 14 d or until discharge	Capsules containing lactose	Velocities in MCA on TCD >120 cm/s with Lindegaard ratio >3
Jaschinski (2008) (7)	98 (40/58)	Patients (20-80 years old) with aSAH	Pravastatin (40 mg/d) within 24 h after the ictus	Nonstatin treatment	No reported
Lynch (2005) (13)	39 (19/20)	Patients presenting within 48 h of aSAH; clipping/ coiling was performed in 9/10 in the statin group and 8/12 in the placebo group	Simvastatin 80 mg/d for 14 d	Placebo for 14 d	TCD velocities >160 cm/s
Chou (2008) (2)	39 (19/20)	Adults (aged >18 years) with Fisher grade 3 SAH were included; clipping was performed in 17 and 16 patients in the statin group and placebo group, respectively	Simvastatin 80 mg/d until discharge from neurointensive care unit, or ≤21 d	Placebo treatment	CVS on TCD was defined as any peak systolic MCA velocity >200 cm/s and Lindegaard ratio >3
Vergouwen (2009) (25)	32 (16/16)	Patients with symptoms and signs of SAH; clipping was performed in 7 patients and coiling in 24 patients	Capsules containing simvastatin 80 mg/d until d 14 after SAH	Capsules containing placebo	TCD velocities >120 cm/s
Macedo (2009) (14)	21 (11/10)	Patients with aSAH	Simvastatin 80 mg/d	Nonstatin treatment	Not reported
Garg (2013) (4)	38 (19/19)	All patients (>18 years old) underwent surgical clipping with evidence of aSAH	Simvastatin 80 mg/d for 14 d	Placebo for 14 d	TCD velocities >160 cm/s

CVS: cerebral vasospasm, SAH: subarachnoid hemorrhage, MCA: middle cerebral artery, TCD: transcranial Doppler, aSAH: aneurysmal SAH.

#### **META-ANALYSIS RESULTS**

#### **Primary end Points**

Six studies (2, 4, 13, 14, 23, 25) reported CVS on TCD at followup. The aggregated results of these studies suggested that statins did not reduce CVS (*RR*=0.80; 95% CI, 0.53-1.21; *P*=0.29) (Figure 3), and the statistical test result for heterogeneity was moderate (*I*<sup>2</sup>=52%, *p*=0.06). Six studies (2, 4, 7, 13, 23, 25) reported DIND; their results indicated that statins reduced DIND (RR=0.56; 95% CI, 0.41-0.75; p=0.0001) (Figure 4), and the statistical test result for heterogeneity was low ( $l^2=16\%$ , p=0.31). Clinical outcome was reported in 6 studies (2, 4, 7, 14, 23, 25); in these studies, statins did not improve outcomes (RR=0.92; 95% Cl, 0.71-1.20; p=0.54) (Figure 5), and the test result for heterogeneity was low ( $l^2=0\%$ , p=0.81), but results demonstrated that statins were able to reduce mortality in 6 studies (2, 4, 7, 14, 23, 25) (*RR*=0.54; 95% Cl, 0.32-0.91; *p*=0.02) (Figure 6), and for this end point, the test for heterogeneity was low ( $l^2 = 13\%$ , p = 0.33).

#### Secondary end points

Owing to the lack of data, most study results could not be

quantitatively analyzed. Thus, descriptive results were as follows: time to CVS from ictus in the statin group was similar to that of the placebo group. For example, Tseng et al. (23) reported the same average time of CVS onset from ictus as 5.2 days in both groups; Chou et al. showed that time to MCA velocity >200 cm/s was  $5.9\pm2.0$  days in the simvastatin group and  $4.8\pm1.4$  days in the placebo group (2). Tseng et al. reported time to DIND from ictus as 5.6 days, and Chou et al. reported  $6.2\pm2.6$  days in the simvastatin group and  $5.4\pm1.9$ days in the placebo group (2, 23). Vergouwen et al. reported that signs of DIND started at a median of 6 days, similar in both treatment groups (25). Overall, statins did not seem to delay time to CVS and DIND from ictus.

The duration of CVS was reported in 2 trials. Tseng et al. described statins as absolutely shortening the duration of CVS by 1.2 days (23). Chou et al. reported a median duration of CVS of 4 days in the simvastatin group and 1 day in the placebo group (2). Thus, statins might be able to shorten the duration of CVS in these patients.

The adverse events reported in 3 studies (2, 4, 13) showed no significant changes in the level of creatine kinase/aspartate





Figure 1: Risk-of-bias summary.

Figure 2: Risk-of-bias graph.

aminotransferase/alanine aminotransferase (RR=1.90; 95% Cl, 0.55–6.50; p=0.31) (Figure 7), and for these studies, the test result for heterogeneity was low ( $l^2$ =0%, p=0.87).

#### DISCUSSION

Three topics are currently being studied that relate to CVS after

aSAH: endothelin type A receptor antagonists, 3-hydroxy-3methylglutaryl coenzyme A (HMG-CoA) reductase inhibitors for treatment, and cortical spreading depolarization as a mechanism of CVS (21). In this analysis, we focused on statins, a potential addition to the treatment for aSAH, because they are effective for lowering serum cholesterol and have

	Statin	15	Place	bo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% CI
Tseng 2005	17	40	25	40	25.8%	0.68 [0.44, 1.05]	
Lynch 2005	5	19	12	20	14.4%	0.44 [0.19, 1.01]	
Chou 2008	13	19	10	20	22.4%	1.37 [0.80, 2.33]	
Vergouwen 2009	12	16	11	16	25.8%	1.09 [0.71, 1.69]	-
Macedo 2009	1	11	4	10	3.7%	0.23 [0.03, 1.71]	
Garg 2013	3	19	5	19	7.9%	0.60 [0.17, 2.16]	
Total (95% CI)		124		125	100.0%	0.80 [0.53, 1.21]	+
Total events	51		67				
Heterogeneity: Tau <sup>2</sup> =	0.12; Chi2	= 10.5	0, df = 5 (	P = 0.0	)6); l <sup>2</sup> = 52	2%	
Test for overall effect:	Favours Statins Favours Placebo						

Figure 3: Forest plot of the effect of statins versus placebo on CVS, as seen on TCD.

	Statin	IS	Place	bo		<b>Risk Ratio</b>	R	isk Ratio	>	
Study or Subgroup	<b>Events Total</b>		Events	Total	Weight	M-H, Fixed, 95% C	M-H, Fixed, 95% Cl			
Tseng 2005	2	40	12	40	15.2%	0.17 [0.04, 0.70]	87 <b></b>	-		
Jaschinski 2008	15	40	35	58	36.1%	0.62 [0.40, 0.98]		-		
Lynch 2005	5	19	12	20	14.8%	0.44 [0.19, 1.01]				
Chou 2008	7	19	10	20	12.3%	0.74 [0.35, 1.54]	-	-		
Vergouwen 2009	6	16	5	16	6.3%	1.20 [0.46, 3.15]		-		
Garg 2013	5	19	8	19	10.1%	0.63 [0.25, 1.57]	-			
Macedo 2009	1	11	4	10	5.3%	0.23 [0.03, 1.71]		-		
Total (95% CI)		164		183	100.0%	0.56 [0.41, 0.75]		•		
Total events	41		86							
Heterogeneity: Chi <sup>2</sup> =	7.10, df =	6 (P = 0	0.31); l <sup>2</sup> =	16%					10	-+
Test for overall effect: Z = 3.86 (P = 0.0001)							Favours Stat	ins Fav	ours Pla	cebo

Figure 4: Forest plot of the effect of statins versus placebo on DIND.



Figure 5: Forest plot of the effect of statins versus placebo on outcomes.

	Statin	15	Place	00		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% C	M-H. Fixed, 95% Cl
Tseng 2005	2	40	8	40	24.0%	0.25 [0.06, 1.11]	
Jaschinski 2008	9	40	13	58	31.9%	1.00 [0.47, 2.12]	+
Chou 2008	0	19	3	20	10.3%	0.15 [0.01, 2.72]	
Vergouwen 2009	2	16	2	16	6.0%	1.00 [0.16, 6.25]	
Macedo 2009	2	11	6	10	18.9%	0.30 [0.08, 1.17]	
Garg 2013	1	19	3	19	9.0%	0.33 [0.04, 2.93]	
Total (95% CI)		145		163	100.0%	0.54 [0.32, 0.91]	•
Total events	16		35				
Heterogeneity: Chi <sup>2</sup> = {	5.73, df =	5 (P = (	0.33); l <sup>2</sup> =	13%			
Test for overall effect:	Z = 2.31 (	P = 0.0	2)				Favours Statins Favours Placebo

Figure 6: Forest plot of the effect of statins versus placebo on mortality.

	Statin	s	Place	00		<b>Risk Ratio</b>	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% C	M-H, Fixed, 95% Cl
Lynch 2005	1	19	0	20	14.1%	3.15 [0.14, 72.88]	
Chou 2008	1	19	1	20	28.1%	1.05 [0.07, 15.66]	
Garg 2013	4	19	2	19	57.8%	2.00 [0.41, 9.65]	
Total (95% CI)		57		59	100.0%	1.90 [0.55, 6.50]	+
Total events	6		3				
Heterogeneity: Chi <sup>2</sup> =	0.29, df = 2	2 (P = (	0.87); l <sup>2</sup> =	0%			
Test for overall effect:	Z = 1.02 (F	= 0.3	1)				Favours Statins Favours Placebo

Figure 7: Forest plot of the effect of statins versus placebo on CK/AST/ALT elevation.

antioxidant, anti-inflammatory, vasodilatory, angiogenic, neuroprotective, vasoprotective, and endothelial preserving effects (19).

Two previous meta-analyses (22, 24) focused on statins in 2008 and 2010; however, their results were conflicting. The data in Sillberg et al. (22) supported "the routine use of statins in the care of patients with aSAH" by calculating the odds ratios (*ORs*) of CVS, DIND, and mortality from 3 RCTs (2, 13, 23), whose results were considered to be subjective (3, 10). In contrast, Vergouwen et al. (24) included 4 RCTs (2, 13, 23, 25) whose *RRs* did not lend statistically significant support to the finding of a beneficial effect of statins in patients with aSAH. In addition, 3 trials addressed this topic during 2009 to 2013, which were necessary to include in another meta-analysis, but Jaschinski et al. (7) and Macedo et al. (14) published only the abstracts, and these had lower methodological quality than comparable studies.

The results of this meta-analysis show that statins are comparable with placebo for reducing DIND and mortality, and these findings are in agreement with those of Sillberg et al. (22). However, these findings show that statins do not reduce CVS as seen on TCD, or improve patient outcomes, similar to the findings of Vergouwen et al. (24). Unlike Sillberg et al. (22) and Vergouwen et al. (24), we tried to compare the

methodological quality not report on the definition (7, 14). Because it differs from a more objective indicator, CVS on TCD could have influenced the judgment of patient outcomes. In addition, we found that

other positive drugs affected the results in the process. Chou et al. (2) administered an endovascular intervention for CVS, and intra-arterial nicardipine in the statin and placebo groups. All patients in Vergouwen et al. (25) received standard care, including treatment with nimodipine 360 mg/day orally in 2 groups. Owing to differences in definitions and positive drugs

time to CVS and DIND from ictus and duration of CVS, but

our results were not clear because few data were available to

draw firm conclusions; the available data demonstrated that statins tended to shorten the duration of CVS, but not delay

The conclusions of our analysis should be approached with

caution for the following reasons: first, many sources of

heterogeneity affected the results, and we found different

definitions of end points and interference from positive

drugs among the studies that may have contributed to

statistical heterogeneity and clinical heterogeneity. Two

RCTs defined CVS on TCD as mean flow velocities >120 cm/s

with a Lindegaard ratio (LR) >3 in MCA (23, 25); another

2 RCTs defined CVS as velocities >160 cm/s (4, 13); 1 RCT

chose velocities >200 cm/s with LR >3 (2); and 2 RCTs did

the time to CVS and DIND from ictus.

used among the studies, the heterogeneity of CVS on TCD was great ( $l^2$ =52%, p=0.06). When effect size was changed to OR, results showed no significant difference (OR=0.59; 95% Cl, 0.28-1.25; p=0.17), and the test result for heterogeneity was low ( $l^2$ =40%, p=0.14). When DIND was considered, definitions used by the RCTs were similar, defined as any 2 or more point folds in the modified Glasgow Coma Scale. Statins were able to reduce DIND by 20% compared with placebo, and the result was relatively stable when excluding lower-quality studies or using a random-effects model. When effect size was changed to OR, data showed similar results (OR=0.38; 95% CI, 0.24-0.61; p<0.0001), and the test result for heterogeneity was low ( $l^2=6\%$ , p=0.38). A poor outcome was defined as "dead," "persistent vegetative state," and "severe disability" in the Glasgow Outcome Scale (GOS) or Modified Rankin Scale (MRS), which was similar among the studies, guantitative results stably showing no difference between the 2 groups. These result were similar to previous RR results when the effect size was changed to OR (OR=0.86; 95% CI, 0.52-1.40; p=0.54), and the test result for heterogeneity was low ( $l^2=0\%$ , p=0.82). As for mortality, our meta-analysis showed that statins decrease it by 10% compared with placebo, but it might be an inconsistent effect based on the length of follow-up and the selection of positive drugs; when OR was used as effect size, the result was OR=0.47; 95% CI, 0.25-0.89 (p=0.02), and the test result for heterogeneity was low ( $l^2=16\%$ , p=0.31); when a random-effects model was used, the effect on mortality was no longer sufficiently robust to remain statistically significant (RR=0.55; 95% Cl, 0.30-1.03; p=0.06). Thus, we suspect that statins play a positive role in improving this outcome.

By comparing data on these outcomes, statins were shown to possibly reduce DIND and mortality, but not CVS on TCD and the likelihood of poor outcomes. We suspect that the choice of acute statins treatment for aSAH is not definitely beneficial, and our results reveal that statins' role is finite. We cannot prove this hypothesis at this stage, however, because relevant data are lacking.

### **Study Limitations**

This meta-analysis had several potential limitations that should be taken into account. First, statins in this metaanalysis involved pravastatin and simvastatin, which differ from each other (19). We did not analyze them in subgroups because of their few statistical differences and heterogeneity. Second, follow-up was different, as Garg et al. (4) reported the functional outcomes at 1, 3 and 6 months after hospital discharge, while other studies recorded outcomes until patient discharge, which was too short to compare the effectiveness. Third, 2 studies were of poor methodological quality (7, 14), and 4 studies were not detailed enough to report randomsequence generation and allocation concealment (2, 7, 13, 14). In addition, we could not obtain original data from articles even though we tried to connect with authors, and 2 trials (7, 14) reported results without particularly detailed information.

### CONCLUSION

Preventing CVS is of great importance. Studies have indicated an increased risk of CVS in patients who are treated with endovascular coiling (17), although surgical clipping of a ruptured aneurysm has not been described as increasing the risk of subsequent CVS (6, 18). In the RCTs included in this analysis, patients in 4 trials underwent coiling or clipping procedures (2, 13, 23, 25), and all the patients in 1 trial (4) underwent surgical clipping of an aneurysm. Thus, aSAH and coiling were common reasons for CVS, which might also have contributed to the heterogeneity of these studies. These intriguing clues may be useful for directing future research on this topic.

Further studies should focus on the following points. First, there is a need to report detailed information according to the Consolidated Standards of Reporting Trials and to intentionally enroll patients undergoing clipping or coiling. Second, the definitions of end points should be unified, just as the definition of CVS should be standardized. Third, although statins are generally considered safe and well tolerated, future studies should pay more attention to their safety, especially with regard to rhabdomyolysis and hepatic function.

In conclusion, current limited evidence suggests that statins might be able to decrease DIND and mortality in aSAH, but not CVS on TCD or the risk of poor outcomes. Thus, statins could play a limited role, and acute statins treatment might be not a good choice for CVS after aSAH. Despite these mixed findings, the results of this analysis should be interpreted with caution owing to the heterogeneity among the study designs it included. Further large-scale, well-designed RCTs on this topic are still needed. A RCT of statin therapy for aSAH, expected to enroll 1600 patients, may clarify those answers (15).

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